Tauberian conditions for Statistical Convergence from Statistical Summability using Nörlund-Euler means in Non-Archimedean Fields

Eunice Jemima1, Srinivasan Vaithinathasamy²

¹Department of Mathematics, Saveetha Engineering College, Thandalam, Chennai- 602105, India, Email: eunicejem@gmail.com ²Professor of Mathematics(Retd.), Email: drvsrinivas.5@gmail.com

ABSTRACT

All through this research article, K stands for a non-archimedean field which is both non-trivially valued and complete. Tauberian conditions for statistical convergence that arises from the fact of statistical summability by Nörlund method over K has recently been investigated by us. We present here, the if and only if conditions for statistical convergence due to statistical summability by Nörlund-Euler methods over non-archimedean fields.

Keywords: Non-archimedean fields, Nörlund-Euler means, statistical convergence, statistical summability $(N, p, q)(E, 1)$ method, Tauberian conditions.

1. INTRODUCTION

It was Fast [4] who first brought about this concept of statistical convergence in the year 1951, which is very significant in summability theory. Schoenberg established a relation between summability and statistical convergence. Hardy's and Landau's Tauberian theorems in classical nature of statistical convergence were studied by Fridy and Khan. In this paper, investigation for Tauberian conditions for sequences that are statistically $(N, p, q)(E, 1)$ summable by Nörlund-Euler means over K is done. Statistical convergence of a sequence (x_k) , is defined as

lim $\frac{1}{n}$ →∞ $\frac{1}{n}$ $\frac{1}{n} |\{k \leq n : |x_k - L| \geq \epsilon\}| = 0,$

 $F\text{or}x_k \in K$, $k = 1,2,...$, for every $\epsilon > 0$. Here L is the statistical limit and the outer vertical lines stand for the set's cardinal number (see [12]). This is written as,

$$
st - \lim_{k \to \infty} x_k = L \# (1.1)
$$

Consider two sequences $p = (p_n)$ and $q = (q_n)$ in K, with

$$
P_n = \sum_{i=0}^{n} p_i
$$
, $p_i \neq 0$ and $Q_n = \sum_{i=0}^{n} q_i$, $q_i \neq 0$

The convolution of (p_n) and (q_n) is given by

$$
R_n = \sum_{k=0} p_k q_{n-k}, \qquad n = 0, 1, 2, ...
$$

Definition 1.1. The summability of $x = (x_n)$ by the generalized Nörlund method (N, p, q) , by (p_n) and (q_n) to s is defined as

$$
t_n^{p,q} = \frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} x_k \to s, \text{ as } n \to \infty
$$

Definition 1.2. A sequence (x_n) is $(E, 1)$ summable to s if,

$$
E_n^1 = \frac{1}{2^n} \sum_{k=0}^n {n \choose k} x_k \to s, \text{ as } n \to \infty
$$

=0 **Definition 1.3.** The Nörlund-Euler method is defined as

$$
t_n^{p,q,E} = \frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} E_k^1
$$

=
$$
\frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} x_j
$$

Definition 1.4. A sequence (x_n) is summable by the generalized Nörlund-Euler $(N, p, q)(E, 1)$ method to *s*, if, $t_n^{p,q,E} \to s$ as $n \to \infty$.

We now define $(N, p, q)(E, 1)$ statistical summability of a sequence as follows.

Definition 1.5. $x = (x_n)$ is said to be statistically $(N, p, q)(E, 1)$ summable to a limit *L* if,

$$
st - \lim_{n \to \infty} t_n^{p,q,E} = L \# (1.2)
$$

2. Key Results

Theorem 2.1. Consider sequences (p_m) and (q_m) in K such that $p_0 > 0$, $p_m \neq 0$, $q_0 > 0$, $q_m \neq 0$. Consider (λ_k) , $\lambda_k \in K$ such that $\lim_{k \to \infty} \lambda_k = 0$ and

$$
\text{st} - \lim_{m \to \infty} \frac{R_m}{R_{\lambda_m}} < 1 \quad \text{for every} \quad 0 < \lambda_m < 1 \# (2.1)
$$

If (x_k) is statistically $(N, p, q)(E, 1)$ summable to a limit L, then (x_k) is said to be statistically convergent to L if and only if,

$$
st - \lim_{m \to \infty} \frac{1}{(R_m - R_{\lambda_m})} \sum_{k = \lambda_m + 1}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) = 0 \# (2.2)
$$

To prove this theorem, we require the lemmas given below.

Lemma 2.1. Consider sequences (p_m) and (q_m) in K such that $p_0 > 0$, $p_m \neq 0$, $q_0 > 0$, $q_m \neq 0$; and

$$
st - \lim_{m \to \infty} \frac{R_m}{R_{\lambda_m}} < 1 \quad \text{for every} \quad 0 < \lambda_m < 1
$$

Let (x_k) be $(N, p, q)(E, 1)$ statistically summable to L. Then, for each $0 < \lambda_m < 1$, $st - \lim_{m \to \infty} t_{\lambda_m}^{p,q,E} = L#(2.3)$

where $\left(R_m \right)$ and $\left(t_{\lambda_m} \right)$ are non-decreasing sequences of positive numbers.

Proof. Given that the sequence (x_k) is statistically summable $(N, p, q)(E, 1)$ to L. i.e., $st - \lim_{m \to \infty} t_m^{p,q,E} = L$ i.e., $\lim_{M \to \infty}$ 1 $\frac{1}{M}\left|\left\{m\leq M;\left|t_m^{p,q,E}-L\right|\geq \epsilon\right\}\right|=0$ i.e., $\lim_{M \to \infty}$ 1 $\frac{1}{M}\Big|\Big\{m\leq M\colon \Big|\frac{1}{R_n}\Big|$ $\frac{1}{R_m} \sum_{k=0}^m p_k q_{m-k} \frac{1}{2^k}$ $\frac{1}{2^k} \sum_{j=0}^k \binom{k}{j}$ $\left| \binom{\kappa}{j} x_j - L \right| \geq \epsilon$ = 0(2.4)

To prove, st $-\lim_{m\to\infty}t_{\lambda_m}^{p,q,E}=L$,

That is, to prove
\n
$$
\lim_{M \to \infty} \frac{1}{M} \left\{ \lambda_m \le M : \left| \frac{1}{R_{\lambda_m}} \sum_{k=0}^{\lambda_m} p_k q_{\lambda_m - k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} x_j - L \right| \ge \epsilon \right\} = 0,
$$
\nconsider
\n
$$
\frac{1}{M} \left\{ \lambda_m \le M : \left| \frac{1}{R_{\lambda_m}} \sum_{k=0}^{\lambda_m} p_k q_{\lambda_m - k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} x_j - L \right| \ge \epsilon \right\}
$$
\n
$$
= \frac{1}{M} \left\{ \lambda_m \le M : \left| \frac{R_m}{R_{\lambda_m}} \sum_{k=0}^{\lambda_m} p_k q_{\lambda_m - k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} x_j - L \right| \ge \epsilon \right\}
$$
\n
$$
\le \frac{1}{M} \left\{ \left\{ m \le M : \left| \frac{1}{R_m} \sum_{k=0}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} x_j - L \right| \ge \epsilon \right\} \right\} \xrightarrow{0 \text{ as } M \to \infty} \text{ (using (2.4))}
$$
\nTherefore,
\n
$$
\lim_{M \to \infty} \frac{1}{M} \left\{ \lambda_m \le M : \left| \frac{1}{R_{\lambda_m}} \sum_{k=0}^{\lambda_m} p_k q_{\lambda_m - k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} x_j - L \right| \ge \epsilon \right\} = 0
$$

$$
f_{\rm{max}}(x)
$$

 $j=0$

 $k=0$

That is, $st - \lim_{m \to \infty} t_{\lambda_m}^{p,q,E} = L$, which proves the lemma.

Lemma 2.2. For $0 < \lambda_m < 1$,

$$
\frac{1}{(R_m - R_{\lambda_m})} \sum_{k=\lambda_m+1}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} x_j
$$

= $t_m^{p,q,E} + \frac{R_{\lambda_m}}{R_m - R_{\lambda_m}} \left(t_m^{p,q,E} - t_{\lambda_m}^{p,q,E} \right)$

provided $R_m > R_{\lambda_m}$.

Proof. Consider the right hand side.

$$
t_{m}^{p,q,E} + \frac{R_{\lambda_{m}}}{R_{m} - R_{\lambda_{m}}} \left(t_{m}^{p,q,E} - t_{\lambda_{m}}^{p,q,E} \right)
$$

\n
$$
= \frac{R_{m} t_{m}^{p,q,E} - R_{\lambda_{m}} t_{m}^{p,q,E} + R_{\lambda_{m}} t_{m}^{p,q,E} - R_{\lambda_{m}} t_{\lambda_{m}}^{p,q,E}}{R_{m} - R_{\lambda_{m}}}
$$

\n
$$
= \frac{1}{R_{m} - R_{\lambda_{m}}} \left[R_{m} \left(\frac{1}{R_{m}} \sum_{k=0}^{m} p_{k} q_{m-k} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} x_{j} \right) - R_{\lambda_{m}} \left(\frac{1}{R_{\lambda_{m}}} \sum_{k=0}^{k} p_{k} q_{\lambda_{m-k}} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} x_{j} \right) \right]
$$

\n
$$
= \frac{1}{R_{m} - R_{\lambda_{m}}} \left[\sum_{k=0}^{\lambda_{m}} p_{k} q_{\lambda_{m-k}} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} x_{j} + \sum_{k=0}^{m} p_{k} q_{\lambda_{m-k}} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} x_{j} \right]
$$

\n
$$
+ \sum_{k=\lambda_{m}+1}^{m} p_{k} q_{m-k} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} x_{j} - \sum_{k=0}^{\lambda_{m}} p_{k} q_{\lambda_{m-k}} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} x_{j}
$$

\n
$$
= \frac{1}{(R_{m} - R_{\lambda_{m}})} \sum_{k=\lambda_{m}+1}^{m} p_{k} q_{m-k} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} x_{j}
$$

Thus,

$$
\frac{1}{(R_m - R_{\lambda_m})} \sum_{k=\lambda_m+1}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} x_j
$$

= $t_m^{p,q,E} + \frac{R_{\lambda_m}}{(R_m - R_{\lambda_m})} (t_m^{p,q,E} - t_{\lambda_m}^{p,q,E})$

proving the lemma. Rearrangement of terms and addition of x_m results in,

$$
x_{m} - t_{m}^{p,q,E} = \frac{R_{\lambda_{m}}}{(R_{m} - R_{\lambda_{m}})} \left(t_{m}^{p,q,E} - t_{\lambda_{m}}^{p,q,E} \right)
$$

$$
- \frac{1}{(R_{m} - R_{\lambda_{m}})} \sum_{k=\lambda_{m}+1}^{m} p_{k} q_{m-k} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} x_{j} + x_{m}
$$

$$
= \frac{R_{\lambda_{m}}}{(R_{m} - R_{\lambda_{m}})} \left(t_{m}^{p,q,E} - t_{\lambda_{m}}^{p,q,E} \right) + \frac{1}{(R_{m} - R_{\lambda_{m}})} \sum_{k=\lambda_{m}+1}^{m} p_{k} q_{m-k} \frac{1}{2^{k}} \sum_{j=0}^{k} {k \choose j} (x_{m} - x_{j})
$$
(2.5)

Proof of the theorem.

Necessity: Under the condition that (x_m) is statistically $(N, p, q)(E, 1)$ summable to a limit L, we first assume*st*−lim_{*m*→∞} $x_m = L$ to prove that, for each 0 < λ_m < 1,

$$
st - \lim_{m \to \infty} \frac{1}{(R_m - R_{\lambda_m})} \sum_{k=\lambda_m+1}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) = 0
$$

i.e., we haves $t-\lim_{m\to\infty}t_m^{p,q,E}=L$ and $\,st-\lim_{m\to\infty}x_m=L$, which implies that $st - \lim_{m \to \infty} (x_m - t_m^{p,q,E}) = 0$

1

i.e.,

That is,

$$
\lim_{M \to \infty} \frac{1}{M} |\{m \le M : |x_m - t_m^{p,q,E}| \ge \epsilon\}| = 0
$$
\n
$$
0 = \lim_{M \to \infty} \frac{1}{M} \left| \left\{ m \le M : \left| \frac{R_{\lambda_m}}{(R_m - R_{\lambda_m})} \left(t_m^{p,q,E} - t_{\lambda_m}^{p,q,E} \right) \right| \right\} + \frac{1}{(R_m - R_{\lambda_m})} \sum_{k = \lambda_m + 1}^{m} p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) \right| \ge \epsilon \left| \left| \frac{1}{(R_m - R_{\lambda_m})} \left(t_m^{p,q,E} - t_{\lambda_m}^{p,q,E} \right) \right| \ge \epsilon \right|
$$
\n
$$
= \lim_{M \to \infty} \max \left\{ \frac{1}{M} \left| \left\{ m \le M : \left| \frac{R_{\lambda_m}}{(R_m - R_{\lambda_m})} \sum_{k = \lambda_m + 1}^{m} p_k q_{m-k} \right. \right\} \right| \ge \epsilon \right\} \right|,
$$
\n
$$
\frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) \left| \ge \epsilon \right\} \left| \right\}
$$
\n
$$
= \lim_{M \to \infty} \frac{1}{M} \left| \left\{ m \le M : \left| \frac{1}{(R_m - R_{\lambda_m})} \sum_{k = \lambda_m + 1}^{m} p_k q_{m-k} \right. \right\}
$$
\n
$$
\frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) \right| \ge \epsilon \left\} \left| \frac{1}{(R_m - R_{\lambda_m})} \sum_{k = \lambda_m + 1}^{m} p_k q_{m-k}
$$
\n
$$
\frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) \right| \ge \epsilon \left\} \left| \frac{1}{(R_m - R_{\lambda_m})} \sum_{k=1}^{m} p_k q_{m-k} \right|
$$
\n
$$
= \lim_{M \to \infty} \frac{1}{M} \left| \left\{ m \le M : \left| \frac{1}{(
$$

since, by (1.2) and (2.3) we have, 1 $\frac{1}{M}\left|\left\{m\leq M:\left|\frac{R_{\lambda_m}}{(R_m-R)}\right|\right|\right\}$ $\frac{R_{\lambda_m}}{(R_m-R_{\lambda_m})}\Bigl(t^{p,q,E}_m-t^{p,q}_{\lambda_m}$ $\left| \begin{matrix} p,q,E \\ \lambda_m \end{matrix} \right| \geq \epsilon \left| \begin{matrix} \end{matrix} \right| \to 0 \quad \text{as} \quad M \to \infty.$ Thus,

$$
st - \lim_{m \to \infty} \frac{1}{(R_m - R_{\lambda_m})} \sum_{k=\lambda_m+1}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) = 0
$$

Sufficiency: Here, we assume that

$$
st - \lim_{m \to \infty} \frac{1}{(R_m - R_{\lambda_m})} \sum_{k=\lambda_m+1}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) = 0
$$

and prove that

$$
st-\lim_{m\to\infty}x_m=L.
$$

For this, we need to prove that

$$
st-\lim_{m\to\infty}\left(x_m-t_m^{p,q,E}\right)=0.
$$

i.e., to prove

$$
\lim_{M\to\infty}\frac{1}{M}\left|\left\{m\leq M\colon\left|x_m-t_m^{p,q,E}\right|\geq\epsilon\right\}\right|=0.
$$

Using equation (2.5),

$$
\frac{1}{M} \left| \{ m \le M : \left| x_m - t_m^{p,q,E} \right| \ge \epsilon \} \right|
$$
\n
$$
= \frac{1}{M} \left| \left\{ m \le M : \left| \frac{R_{\lambda_m}}{(R_m - R_{\lambda_m})} \left(t_m^{p,q,E} - t_{\lambda_m}^{p,q,E} \right) \right| + \frac{1}{(R_m - R_{\lambda_m})} \sum_{k=\lambda_{m+1}}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) \right| \ge \epsilon \right\} |
$$
\n
$$
\le \max \left\{ \frac{1}{M} \left| \left\{ m \le M : \left| \frac{R_{\lambda_m}}{(R_m - R_{\lambda_m})} \left(t_m^{p,q,E} - t_{\lambda_m}^{p,q,E} \right) \right| \ge \epsilon \right\} \right|,
$$
\n
$$
\frac{1}{M} \left| \left\{ m \le M : \left| \frac{1}{(R_m - R_{\lambda_m})} \sum_{k=\lambda_{m+1}}^m p_k q_{m-k} \right| \right\} - \frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) \right| \ge \epsilon \right\} |
$$

By our assumption,

$$
\frac{1}{M} \left| \left\{ m \le M : \left| \frac{1}{(R_m - R_{\lambda_m})} \sum_{k=\lambda_m+1}^m p_k q_{m-k} \frac{1}{2^k} \sum_{j=0}^k {k \choose j} (x_m - x_j) \right| \ge \epsilon \right\} \right|
$$

\n
$$
\rightarrow 0 \text{ as } M \rightarrow \infty
$$

Therefore,

$$
\frac{1}{M} \left| \{ m \le M : |x_m - t_m^{p,q,E}| \ge \epsilon \} \right|
$$
\n
$$
\le \max \left\{ \frac{1}{M} \left| \left\{ m \le M : \left| \frac{R_{\lambda_m}}{(R_m - R_{\lambda_m})} \left(t_m^{p,q,E} - t_{\lambda_m}^{p,q,E} \right) \right| \ge \epsilon \right\} \right|, 0 \right\}
$$
\n
$$
\le \frac{1}{M} \left| \left\{ m \le M : \left| \frac{R_{\lambda_m}}{(R_m - R_{\lambda_m})} \left(t_m^{p,q,E} - t_{\lambda_m}^{p,q,E} \right) \right| \ge \epsilon \right\} \right|
$$
\n
$$
\to 0 \text{ as } M \to \infty \text{ by (1.2) and (2.3)},
$$

by which we have,

$$
\lim_{M\to\infty}\frac{1}{M}\left|\{m\leq M:\left|x_m-t_m^{p,q,E}\right|\geq\epsilon\}\right|=0
$$

That is,

$$
st-\lim_{m\to\infty}\left(x_m-t_m^{p,q,E}\right)=0
$$

Thus the theorem is proved.

Acknowledgment

I would like to extend my sincere gratitude to my Research supervisor Prof. Srinivasan Vaithinathasamy for his valuable guidance. With his knowledge and expertise I can delve into research connecting non-Archimedean analysis with fuzzy math as seen in [13] and [14].

REFERENCES

- [1] E. Aljimi, E. Hoxha, V. Loku, Some results of weighted Nörlund-Euler statistical convergence, International Mathematical Forum 8, No. 37 (2013) 1797-1812.
- [2] G. Bachman: Introduction to p-Adic Numbers and Valuation Theory, Academic Press, 1964.
- [3] N.L. Braha, A Tauberian theorem for the generalized Nörlund-Euler summability method, Journal of inequalities and special functions 7, No. 4 (2016) 137-142.
- [4] D. Eunice Jemima, V. Srinivasan, Nörlund Statistical Convergence and Tauberian conditions for Statistical Convergence from Statistical Summability using Nörlund means in non-Archimedean Fields, Journal of Mathematics and Computer Science 24, (2022) 299-307.
- [5] H. Fast, Sur la convergence statistique, Colloq. Math. 2, (1951) 241244.
- [6] J.A. Fridy, On statistical Convergence, Analysis 5, (1985) 301-313.
- [7] G.H. Hardy: Divergent Series, Oxford Univ. Press, London, 1949.
- [8] Huseyin Cakalli, A Study on statistical convergence, Functional Analysis, Approximation and Computation 1, No. 2 (2009) 19-24.
- [9] A.F. Monna, Sur le theoreme de Banach-Steinhaus, Indag. Math. 25, (1963) 121-131.
- [10] P.N. Natarajan, V. Srinivasan, Convolution of Nörlund methods in non-archimedean fields, Ann. Math. Blaise Pascal 4, No. 2 (1997) 41-47.
- [11] V. Srinivasan, P.N. Natarajan, On generalized Nörlund methods in non-archimedean fields, Bulletin of Allahabad Mathematical Society 15, (2000) 43 − 46.
- [12] K. Suja, V. Srinivasan, On statistically convergent and statistically cauchy sequences in nonarchimedean fields, Journal of Advances in Mathematics 6, No. 3 (2014) 1038-1043.
- [13] K. Kalaiarasi, H. Mary Henrietta, M. Sumathi, Applying Hessian Matrix Techniques to obtain the efficient optimal order quantity using Fuzzy parameters, Communications in Mathematics and Applications. 13 (2): 725 - 735 (2022).
- [14] K. Kalaiarasi, M. Sumathi, H. Mary Henrietta, Optimization of fuzzy inventory model for EOQ using Lagrangian method, Malaya Journal of Matematik 7(3):497-501 (2019).